

# SHORT PROOF OF THE BERNOULLI CONJECTURE

**ABSTRACT.** We give a short proof of the Bernoulli conjecture. The proof was produced by an internal model at OpenAI.

## 1. INTRODUCTION

Let  $\varepsilon = (\varepsilon_i)_{i \in I}$  have independent uniform sign coordinates, and let  $G = (g_i)_{i \in I}$  have independent standard Gaussian coordinates. For a nonempty  $T \subset \ell^2(I)$ , write

$$b(T) = \mathbb{E} \sup_{t \in T} \sum_{i \in I} t_i \varepsilon_i, \quad g(T) = \mathbb{E} \sup_{t \in T} \sum_{i \in I} t_i g_i,$$

with the usual finite-subset interpretation. Recall that  $X \preceq_{\text{cx}} Y$  means  $\mathbb{E}\Phi(X) \leq \mathbb{E}\Phi(Y)$  for every convex  $\Phi$  for which the expectations are defined.

**Theorem 1.1** (Bernoulli conjecture). *There is a universal constant  $C$  such that, whenever  $\emptyset \neq T \subset \ell^2(I)$  and  $b(T) < \infty$ , there exist  $U \subset \ell^1(I)$  and  $V \subset \ell^2(I)$  such that*

$$T \subset U + V, \quad \sup_{u \in U} \|u\|_1 \leq Cb(T), \quad g(V) \leq Cb(T).$$

The conjecture was proposed by Talagrand [9] and first proved by Bednorz and Latała [2]. The main input for our proof is a subgaussian convex-comparison theorem, which follows by duality from Liu’s comparison principle [5, Corollary 2] and was explicitly formulated in convex order by van Handel [10]. In fact, the key input for Theorem 1.1 is the following convex domination result.

**Theorem 1.2** (Gaussian-to-Bernoulli comparison). *There is a universal constant  $C_0$  such that, for every  $G \sim N(0, I_n)$  and every uniform  $\varepsilon \in \{-1, 1\}^n$ ,*

$$X \preceq_{\text{cx}} G, \quad \|X\|_\infty \leq 1 \text{ almost surely} \implies X \preceq_{\text{cx}} C_0 \varepsilon.$$

We use this Gaussian comparison in the following form: for a universal  $C_G$ , if  $Y$  is centered and

$$\log \mathbb{E} e^{\langle a, Y \rangle} \leq \frac{K^2}{2} \|a\|_2^2 \quad (a \in \mathbb{R}^n), \tag{1}$$

then  $Y \preceq_{\text{cx}} C_G K G'$ , where  $G' \sim N(0, I_n)$  [10]. By Strassen’s theorem [8], this is equivalent to the existence of a martingale coupling with

$$Y = C_G K \mathbb{E}[G' \mid Y]. \tag{2}$$

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**Note on concurrent work:** During the preparation of this manuscript, we learned of an independent alternate proof of the Bernoulli conjecture by Liu and Zadik [6].

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## 2. REDUCTION TO THEOREM 1.2

We first give the reduction from Theorem 1.1 to Theorem 1.2.

*Proof.* By the standard finite-section and compactness reduction, it suffices to prove Theorem 1.1 for a finite  $T \subset \mathbb{R}^n$ . Set

$$D(T) = \inf_{(v_t)_{t \in T}} \left\{ \max_{t \in T} \|t - v_t\|_1 + \mathbb{E} \max_{t \in T} \langle v_t, G \rangle \right\}. \quad (3)$$

For a measurable probability kernel  $q = (q_t)_{t \in T}$ , put

$$q_t \geq 0, \quad \sum_{t \in T} q_t = 1, \quad m_t = \mathbb{E}[Gq_t(G)].$$

We claim that

$$D(T) = \sup_q \left\{ \sum_{t \in T} \langle t, m_t \rangle : \sum_{t \in T} \|m_t\|_\infty \leq 1 \right\}. \quad (4)$$

Indeed,

$$\begin{aligned} \max_t \|t - v_t\|_1 &= \sup_{\sum_t \|z_t\|_\infty \leq 1} \sum_t \langle t - v_t, z_t \rangle, \\ \mathbb{E} \max_t \langle v_t, G \rangle &= \sup_q \sum_t \langle v_t, m_t \rangle. \end{aligned}$$

The kernels form a weakly compact convex subset of  $L^2(\gamma_n)^T$ , so one application of minimax gives

$$D(T) = \sup_{\substack{\sum_t \|z_t\|_\infty \leq 1 \\ q_t \geq 0, \sum_t q_t = 1}} \inf_{(v_t)} \left\{ \sum_t \langle t, z_t \rangle + \sum_t \langle v_t, m_t - z_t \rangle \right\}.$$

The inner infimum is finite only when  $z_t = m_t$  for every  $t$ : if  $m_s \neq z_s$ , take  $v_s = -R(m_s - z_s)$  and  $v_t = 0$  for  $t \neq s$  and let  $R \rightarrow \infty$ . In the remaining case the inner expression is  $\sum_t \langle t, m_t \rangle$ . This proves (4) and, in particular, explains precisely why the dual variable is forced to equal the Gaussian barycenter.

Fix an admissible kernel in (4), and write  $p_t = \mathbb{E}q_t(G)$  and  $a_t = \|m_t\|_\infty$ . Since  $\sum_t (a_t - p_t)_+ \leq \sum_t a_t \leq 1$ , choose a probability vector  $r = (r_t)_{t \in T}$  such that  $r_t \geq (a_t - p_t)_+$ . Toss an independent fair coin. On heads, sample a label  $J$  from  $q(G)$ ; on tails, sample  $J$  independently of  $G$  with distribution  $r$ . Then

$$\mathbb{P}(J = t) = \frac{p_t + r_t}{2}, \quad \mathbb{E}[G\mathbf{1}_{\{J=t\}}] = \frac{m_t}{2}.$$

Consequently  $X = \mathbb{E}[G \mid J]$  satisfies  $X \preceq_{\text{cx}} G$ , and on every positive-probability event  $\{J = t\}$ ,

$$X = \frac{m_t}{p_t + r_t}, \quad \|X\|_\infty = \frac{a_t}{p_t + r_t} \leq 1.$$

Apply Theorem 1.2 to the convex function  $\Phi(x) = \max_{t \in T} \langle t, x \rangle$ . We obtain

$$\frac{1}{2} \sum_{t \in T} \langle t, m_t \rangle = \mathbb{E} \langle J, X \rangle \leq \mathbb{E} \Phi(X) \leq C_0 b(T).$$

Taking the supremum in (4) yields  $D(T) \leq 2C_0 b(T)$ . The infimum in (3) is attained by coercivity. For a minimizer  $(v_t)$  set  $U = \{t - v_t : t \in T\}$  and  $V = \{v_t : t \in T\}$ . Both terms in (3) are nonnegative. Hence

$$T \subset U + V, \quad \sup_{u \in U} \|u\|_1 \leq 2C_0 b(T), \quad g(V) \leq 2C_0 b(T),$$

as required.  $\square$

### 3. PROOF OF THEOREM 1.2

The first lemma is the finite-support Gaussian case of a maximum-entropy martingale bridge. It is a special case of the construction of Hua, Song, and Tudose [4, Proposition 3.3 and Section 4.1], applied to  $X/2 \preceq_{\text{cx}} G/2$ . The general theory of martingale Schrödinger bridges is developed by Nutz and Wiesel [7] and by Backhoff, Beiglböck, Bifronte, and Ley [1]. We include the short argument for completeness.

**Lemma 3.1.** *Suppose  $X$  has finite support and  $X \preceq_{\text{cx}} G$ . There exists a coupling of  $X$  and  $G \sim N(0, I_n)$  such that*

$$\mathbb{E}[G \mid X] = \frac{1}{2}X, \quad (5)$$

and each conditional distribution  $G \mid \{X = x\}$  has an everywhere positive, 1-strongly log-concave density. Among couplings satisfying (5), this coupling uniquely maximizes  $H(X \mid G)$ .

*Proof.* Write  $\mathbb{P}(X = x_j) = p_j > 0$  for  $1 \leq j \leq m$ . Convex comparison gives  $\mathbb{E}X = 0$ . Set  $a_m = 0$ ,  $b_m = 0$ , and  $\ell_j(g) = a_j + \langle b_j, g \rangle$ . Consider

$$P(a, b) = \mathbb{E} \log \sum_{j=1}^m e^{\ell_j(G)} - \sum_{j=1}^m p_j \left( a_j + \frac{1}{2} \langle b_j, x_j \rangle \right). \quad (6)$$

Let  $h = \mathbb{E} \max_j \ell_j(G)$ . Since log-sum-exp dominates the maximum,

$$P(a, b) \geq \frac{1}{2} \left( h - \sum_j p_j \ell_j(x_j) \right) + \frac{1}{2} \mathbb{E} \left[ \max_j \ell_j(G) - \sum_j p_j \ell_j(G) \right]. \quad (7)$$

The first term is nonnegative by  $X \preceq_{\text{cx}} G$ . The second is continuous, positively homogeneous, and strictly positive unless all  $\ell_j$  coincide, hence unless  $(a, b) = 0$ . Compactness of the unit sphere now shows that  $P$  is coercive, so it has a minimizer  $(a^*, b^*)$ .

Define

$$q_j(g) = \frac{e^{a_j^* + \langle b_j^*, g \rangle}}{\sum_{k=1}^m e^{a_k^* + \langle b_k^*, g \rangle}}. \quad (8)$$

Differentiating (6) gives

$$\mathbb{E} q_j(G) = p_j, \quad \mathbb{E}[G q_j(G)] = \frac{1}{2} p_j x_j \quad (1 \leq j < m).$$

The identities for  $j = m$  follow by summation from  $\mathbb{E}G = \mathbb{E}X = 0$ . Conditional on  $G = g$ , assign  $X = x_j$  with probability  $q_j(g)$ . This has the prescribed  $X$  marginal and satisfies (5).

To see the maximum-entropy claim, let  $\tilde{q}$  be any other kernel with the same marginal and barycenter constraints, and write  $Z(g) = \sum_k e^{a_k^* + \langle b_k^*, g \rangle}$  and  $H(\tilde{q}) = -\mathbb{E} \sum_j \tilde{q}_j(G) \log \tilde{q}_j(G)$ . Since  $\log q_j(g) = a_j^* + \langle b_j^*, g \rangle - \log Z(g)$ , those constraints give

$$\mathbb{E} \sum_j \tilde{q}_j(G) \log q_j(G) = \sum_j p_j \left( a_j^* + \frac{1}{2} \langle b_j^*, x_j \rangle \right) - \mathbb{E} \log Z(G) = \mathbb{E} \sum_j q_j(G) \log q_j(G).$$

Thus the cross-entropy term is the same for every feasible kernel, and

$$H(q) - H(\tilde{q}) = \mathbb{E} \sum_j \tilde{q}_j(G) \log \frac{\tilde{q}_j(G)}{q_j(G)} \geq 0.$$

Equality holds only when  $\tilde{q} = q$  almost surely. Finally,  $\log q_j$  is concave because  $\log$ -sum-exp is convex. Hence the conditional Gaussian density  $g \mapsto [p_j(2\pi)^{n/2}]^{-1} q_j(g) e^{-\|g\|_2^2/2}$  is everywhere positive and 1-strongly log-concave.  $\square$

The following is an immediate consequence of the Brascamp–Lieb variance inequality [3, Theorem 4.1].

**Lemma 3.2.** *For the coupling in Lemma 3.1,*

$$\text{Var}(G_i | X) \leq 1, \quad \mathbb{E}[G_i^2 | X] \leq 1 + \frac{1}{4} X_i^2.$$

We next record a standard subgaussian estimate needed for Lemma 3.4.

**Lemma 3.3.** *Let  $V_1, \dots, V_n$  be independent and centered, with  $\mathbb{E}e^{|V_i|/\kappa} \leq 2$ , and let  $Y_i = \mathbb{E}[V_i | \mathcal{H}]$  satisfy  $|Y_i| \leq D$ . Then*

$$\log \mathbb{E}e^{\langle a, Y \rangle} \leq \max\{4\kappa^2, 2\kappa D\} \|a\|_2^2 \quad (a \in \mathbb{R}^n).$$

*Proof.* The standard Bernstein moment estimate for a centered subexponential variable [11, proof of Proposition 2.8.1] gives

$$\mathbb{E}e^{sV_i} \leq e^{4\kappa^2 s^2} \quad \text{if } \kappa|s| \leq \frac{1}{2}.$$

Set  $A = \{i : \kappa|a_i| > 1/2\}$ . For  $i \in A$ ,  $D|a_i| \leq 2\kappa D a_i^2$ . Thus conditional Jensen and independence give

$$\mathbb{E}e^{\langle a, Y \rangle} \leq e^{2\kappa D \sum_{i \in A} a_i^2} \prod_{i \notin A} \mathbb{E}e^{a_i V_i} \leq \exp \left( 2\kappa D \sum_{i \in A} a_i^2 + 4\kappa^2 \sum_{i \notin A} a_i^2 \right).$$

The claimed bound follows.  $\square$

**Lemma 3.4.** *Suppose  $\|X\|_\infty \leq 1$ , and use the coupling of Lemma 3.1. For every  $L > 0$ , there exist a uniform  $\varepsilon \in \{-1, 1\}^n$  and a centered  $Y \in \mathbb{R}^n$  such that*

$$\frac{1}{2}X = L\mathbb{E}[\varepsilon | X] + Y, \quad \|Y\|_\infty \leq \frac{1}{2L}, \quad \log \mathbb{E}e^{\langle a, Y \rangle} \leq \frac{4}{L^2} \|a\|_2^2. \quad (9)$$

*Proof.* Independently of  $(G, X)$ , take independent  $U_i \sim \text{Unif}[-L, L]$  and set  $\varepsilon_i = \text{sgn}(G_i + U_i)$ . The pairs  $(G_i, U_i)$  are independent and symmetric. Therefore the  $\varepsilon_i$  are genuinely independent uniform signs; no independence conditional on  $X$  is asserted or needed. Moreover,

$$\mathbb{E}[\varepsilon_i | G, X] = \text{clip}(G_i/L, -1, 1).$$

Define  $V_i = \text{sgn}(G_i)(|G_i| - L)_+$  and  $Y_i = \mathbb{E}[V_i | X]$ . The  $V_i$  are independent and centered, and  $G_i = L \text{clip}(G_i/L, -1, 1) + V_i$ . Taking conditional expectations and using (5) proves the first identity in (9).

The elementary inequality  $(r - L)_+ \leq r^2/(4L)$  gives

$$\mathbb{E}e^{L|V_i|} \leq \mathbb{E}e^{G_i^2/4} = \sqrt{2} \leq 2.$$

Lemma 3.2 and  $|X_i| \leq 1$  further give

$$|Y_i| \leq \frac{\mathbb{E}[G_i^2 | X]}{4L} \leq \frac{1 + X_i^2/4}{4L} \leq \frac{1}{2L}.$$

Apply Lemma 3.3 with  $\kappa = L^{-1}$  and  $D = (2L)^{-1}$  to obtain the final bound in (9).  $\square$

*Proof of Theorem 1.2.* For a convex  $\phi : \mathbb{R}^n \rightarrow \mathbb{R}$ , write

$$S_\phi = \sup\{\mathbb{E}\phi(X) : X \preceq_{\text{cx}} G, \|X\|_\infty \leq 1\} < \infty.$$

It suffices to show that  $S_\phi \leq \mathbb{E}\phi(C_0\varepsilon)$  for every convex  $\phi$ . Conditional expectation on increasingly fine finite partitions of the cube shows that the supremum defining  $S_\phi$  may be taken over finite-valued  $X$ .

Set  $C_* = \max\{2\sqrt{2}C_G, 1/2\}$ , fix  $L > 2C_*$ , and put  $\lambda = C_*/L$  and  $C_0 = 2L/(1 - 2\lambda)$ . For each finite-valued admissible  $X$ , Lemma 3.4 and Gaussian comparison (2) yield a uniform sign vector  $\varepsilon$  and  $G' \sim N(0, I_n)$  such that

$$\frac{1}{2}X = L\mathbb{E}[\varepsilon \mid X] + \lambda Z, \quad Z = \mathbb{E}[G' \mid X].$$

Jensen gives  $Z \preceq_{\text{cx}} G$ , and Lemma 3.4 gives

$$\|Z\|_\infty \leq \frac{1}{2L\lambda} = \frac{1}{2C_*} \leq 1, \quad X = (1 - 2\lambda)\mathbb{E}[C_0\varepsilon \mid X] + 2\lambda Z.$$

Thus another application of Jensen gives

$$\mathbb{E}\phi(X) \leq (1 - 2\lambda)\mathbb{E}\phi(C_0\varepsilon) + 2\lambda\mathbb{E}\phi(Z) \leq (1 - 2\lambda)\mathbb{E}\phi(C_0\varepsilon) + 2\lambda S_\phi.$$

Taking the supremum over  $X$  and using  $2\lambda < 1$  yields  $S_\phi \leq \mathbb{E}\phi(C_0\varepsilon)$ , as claimed.  $\square$

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